# Introduction to Linear Regression 

Arnaud Legrand

MOSIG Lecture

## Outline

(1) Simple Linear Regression

- General Introduction
- Fitting a Line to a Set of Points
(2) Linear Model
- Linear Regression
- Underlying Hypothesis
- Checking hypothesis
- Decomposing the Variance
- Making Predictions
- Confidence interval
(3) Conclusion


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## What is a regression?

Regression analysis is the most widely used statistical tool for understanding relationships among variables. Several possible objectives including:
(1) Prediction of future observations. This includes extrapolation since we all like connecting points by lines when we expect things to be continuous
(2) Assessment of the effect of, or relationship between, explanatory variables on the response

- A general description of data structure (generally expressed in the form of an equation or a model connecting the response or dependent variable and one or more explanatory or predictor variable)
- Defining what you should "expect" as it allows you to define and detect what does not behave as expected
The linear relationship is the most commonly found one
- we will illustrate how it works
- it is very general and is the basis of many more advanced tools (polynomial regression, ANOVA, ...)


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## Starting With a Simple Data Set

Descriptive statistics provides simple summaries about the sample and about the observations that have been made.

How could we summarize the following data set?

|  | $x$ | $y$ |
| ---: | ---: | ---: |
| 1 | 1.00 | 3.00 |
| 2 | 2.00 | 1.00 |
| 3 | 3.00 | 5.00 |
| 4 | 5.00 | 2.00 |
| 5 | 7.00 | 6.00 |
| 6 | 9.00 | 4.00 |
| 7 | 11.00 | 7.00 |
| 8 | 13.00 | 9.00 |
| 9 | 14.00 | 8.00 |
| 10 | 15.00 | 5.00 |



## Eyeball Method



- A straight line drawn through the maximum number of points on a scatter plot balancing about an equal number of points above and below the line
- Some points are rather far from the line. Maybe we should instead try to minimize some kind of distance to the line


## Least Squares Line (1): What to minimize?



Intuitively, a large error is much more important than a small one. We could try to minimize $F(\alpha, \beta)=\sum_{i} \underbrace{\left(y_{i}-\alpha-\beta x_{i}\right)^{2}}_{e_{i}}$, the size of all residuals:

- If they were all zero we would have a perfect line
- Trade-off between moving closer to some points and at the same time moving away from other points


## Least Squares Line (2): Simple Formula

$$
F(\alpha, \beta)=\sum_{i=1}^{n}\left(y_{i}-\alpha-\beta x_{i}\right)^{2}
$$

$F$ is quadratic in $\alpha$ and in $\beta$ so if we simply differentiate $F$ by $\alpha$ and by $\beta$, we can obtain a closed form for the minimum:

$$
\begin{aligned}
\hat{\beta} & =\frac{\sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)\left(y_{i}-\bar{y}\right)}{\sum_{i=1}^{n}\left(x_{i}-\bar{x}\right)^{2}}=\frac{\sum_{i=1}^{n} x_{i} y_{i}-\frac{1}{n} \sum_{i=1}^{n} x_{i} \sum_{j=1}^{n} y_{j}}{\sum_{i=1}^{n}\left(x_{i}^{2}\right)-\frac{1}{n}\left(\sum_{i=1}^{n} x_{i}\right)^{2}} \\
& =\frac{\overline{x y}-\bar{x} \bar{y}}{\overline{x^{2}}-\bar{x}^{2}}=\frac{\operatorname{Cov}[x, y]}{\operatorname{Var}[x]}=r_{x y} \frac{s_{y}}{s_{x}} \\
\hat{\alpha} & =\bar{y}-\hat{\beta} \bar{x}, \text { where: }
\end{aligned}
$$

- $\bar{x}$ and $\bar{y}$ are the sample mean of $x$ and $y$
- $r_{x y}$ is the sample correlation coefficient between $x$ and $y$
- $s_{x}$ and $s_{y}$ are the sample standard deviation of $x$ and $y$

Also has a good geometric interpretation (orthogonal projection)

## Least Squares Line (3): y as a function of $x$ or the opposite?



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## Least Squares Line (3): y as a function of $x$ or the opposite?



OK, do we have less asymetrical options?

## Least Distances Line (a.k.a. Deming Regression)



- Note that somehow, this makes sense only if we have a square plot, i.e., if $x$ and $y$ have the same units


## Least Rectangles Line



- Minimize $E(\alpha, \beta)=\sum_{i=1}^{n}\left|x_{i}-\frac{y_{i}-\alpha}{\beta}\right| \cdot\left|y_{i}-\alpha-\beta x_{i}\right|$
- This leads to the regression line $y=\frac{s_{y}}{s_{x}}(x-\bar{x})+\bar{y}$.


## Least Squares (in Both Directions) Line



- Minimize $D(\alpha, \beta)=\sum_{i=1}^{n}\left(x_{i}-\frac{y_{i}-\alpha}{\beta}\right)^{2}+\left(y_{i}-\alpha-\beta x_{i}\right)^{2}$
- Has to be computed analytically


## Which line to choose?



## What does correspond to each line?

- Eyeball: AFAIK nothing
- Least Squares: classical linear regression $y \sim x$
- Least Squares in both directions: I don't know
- Deming: equivalent to Principal Component Analysis
- Rectangles: may be used when one variable is not "explained" by the other, but are inter-dependent

This is not just a geometric problem. You need a model of to decide which one to use

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## The Simple Linear Regression Model

We need to invest in a probability model

$$
Y=a+b X+\varepsilon
$$

- $Y$ is the response variable
- $X$ is a continuous explanatory variable
- $a$ is the intercept
- $b$ is the slope
- $\varepsilon$ is some noise

- $a+b X$ represents the "true line", the part of $Y$ that depends on $X$
- The error term $\varepsilon$ is independent "idosyncratic noise", i.e., the part of $Y$ not associated with $X$


## The Simple Linear Regression Model

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## Gauss-Markov Theorem

Under a few assumptions, the least squares regression is the best linear unbiased estimate

- $\mathbb{E}(\hat{\beta})=b$ and $\mathbb{E}(\hat{\alpha})=a$
- $\operatorname{Var}(\hat{\beta})$ and $\operatorname{Var}(\hat{\alpha})$ are minimal


## Multiple explanatory variables

- The same results hold true when there are several explanatory variables:

$$
Y=a+b^{(1)} X^{(1)}+b^{(2)} X^{(2)}+b^{(1,2)} X^{(1)} X^{(2)}+\varepsilon
$$

The least squares regressions are good estimators of $a, b^{(1)}, b^{(2)}$, $b^{(1,2)}$

- We can use an arbitrary linear combination of variables, hence

$$
Y=a+b^{(1)} X+b^{(2)} \frac{1}{X}+b^{(3)} X^{3}+\varepsilon
$$

is also a linear model

- Obviously the closed-form formula are much more complicated but softwares like $R$ handle this very well


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## Important Hypothesis (1)

Weak exogeneity The predictor variables $X$ can be treated as fixed values, rather than random variables: the $X$ are assumed to be error-free, i.e., they are not contaminated with measurement errors
Although not realistic in many settings, dropping this assumption leads to significantly more difficult errors-in-variables models
Linearity the mean of the response variable is a linear combination of the parameters (regression coefficients) and the predictor variables Since predictor variables themselves can be arbitrarily transformed, this is not that restrictive. This trick is used, for example, in polynomial regression, but beware of overfitting
Independance of Errors if several responses $Y_{1}$ and $Y_{2}$ are fit, $\varepsilon_{1}$ and $\varepsilon_{2}$ should be independant

## Other Very Important Hypothesis

Constant variance (a.k.a. homoscedasticity)


- Variance is independant of $X$
- If several responses $Y_{1}$ and $Y_{2}$ are fit, $\varepsilon_{1}$ and $\varepsilon_{2}$ should have the same variance
- Either normalize $Y$ or use an other estimator


## Other Classical Hypothesis (3)

Normal and iid errors This is not an assumption of the Gauss Markov Theorem. Yet, it is quite convenient to build confidence intervals of the regression
Arrangement of the predictor variables $X$ it has a major influence on the precision of estimates of $\beta$ (remember Anscombe's quartet).


This is part of your design of experiments:

- If you want to test linearity, $X$ should be uniformly distributed
- If you want the best estimation, you should use extreme values of $X$


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## Linearity: Residuals vs. Explanatory Variable



When there are several factors, you have to check for every dimension. . .

## Homoscedasticity: Residuals vs. Fitted values



## Normality: qqplots



A quantile-quantile plot is a graphical method for comparing two probability distributions by plotting their quantiles against each other

## Model Formulae in R

The structure of a model is specified in the formula like this:

> response variable ~ explanatory variable(s)
~ reads "is modeled as a function of " and $\operatorname{lm}\left(y{ }^{\sim} \mathrm{x}\right)$ means $y=\alpha+\beta x+\varepsilon$
On the right-hand side, on should specify how the explanatory variables are combined. The symbols used here have a different meaning than in arithmetic expressions

-     + indicates a variable inclusion (not an addition)
-     - indicates a variable deletion (not a substraction)
-     * indicates inclusion of variables and their interactions
- : means an interaction

Therefore

- $z^{\sim} \mathrm{x}+\mathrm{y}$ means $\mathrm{z}=\alpha+\beta_{1} x+\beta_{2} y+\varepsilon$
- $z^{\sim}$ x*y means $z=\alpha+\beta_{1} x+\beta_{2} y+\beta_{3} x y+\varepsilon$
- $z^{\sim}(x+y) \wedge 2$ means the same
- $\log (y) \sim I(1 / x)+x+I\left(x^{\wedge} 2\right)$ means $z=\alpha+\beta_{1} \times \frac{1}{x}+\beta_{2} x+\beta_{3} x^{2}+\varepsilon$


## Checking the model with $R$

```
reg <- lm(data=df[df$type=="heteroscedastic",],y^x)
par(mfrow=c(2,2)); plot(reg); par(mfrow=c(1,1))
```



Fitted values


Leverage

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## Decomposing the Variance

How well does the least squares line explain variation in $Y$ ?


## Decomposing the Variance

How well does the least squares line explain variation in $Y$ ?
Remember that $Y=\hat{Y}+\varepsilon \quad(\hat{Y}$ is the "true mean").
Since $\hat{Y}$ and $\varepsilon$ are uncorrelated, we have

$$
\begin{aligned}
\operatorname{Var}(Y) & =\operatorname{Var}(\hat{Y}+\varepsilon)=\operatorname{Var}(\hat{Y})+\operatorname{Var}(\varepsilon) \\
\frac{1}{n-1} \sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2} & =\frac{1}{n-1} \sum_{i=1}^{n}\left(\hat{Y}_{i}-\overline{\hat{Y}}\right)^{2}+\frac{1}{n-1} \sum_{i=1}^{n}\left(\varepsilon_{i}-\bar{\varepsilon}\right)^{2}
\end{aligned}
$$

Since $\bar{\varepsilon}=0_{n}$ and $\bar{Y}=\overline{\hat{Y}}$, we have
$\underbrace{\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}}_{\text {Total Sum of Squares }}=\underbrace{\sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}\right)^{2}}_{\text {Regression SS }}+\underbrace{\sum_{i=1}^{n} \bar{\varepsilon}^{2}}_{\text {Error SS }}$

- $S S R=$ Variation in $Y$ explained by the regression line
- $S S E=$ Variation in $Y$ that is left unexplained

$$
S S R=S S T \Rightarrow \text { perfect fit }
$$

## A Goodness of Fit Measure: $R^{2}$

The coefficient of determination, denoted by $R^{2}$, measures goodness of fit:

$$
R^{2}=\frac{S S R}{S S T}=1-\frac{S S E}{S S T}
$$

- $0 \leqslant R^{2} \leqslant 1$
- The closer $R^{2}$ is to 1 , the better the fit

Warning:

- A not so low $R^{2}$ may mean important noise or bad model
- As you add parameters to a model, you inevitably improve the fit. There is a trade-off beteween model simplicity and fit. Strive for simplicity!


## Illustration with R (homoscedastic data)



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```
reg <- lm(data=df[df$type=="homoscedastic",],y~x)
summary(reg)
```


## Illustration with R (homoscedastic data)

```
reg <- lm(data=df[df$type=="homoscedastic",],y~}\textrm{x}
summary(reg)
```

```
Call:
lm(formula = y ~ x, data = df[df$type == "homoscedastic", ])
```

Residuals:
Min 1Q Median 3Q Max
-4.1248-1.3059 -0.0366 $1.0588 \quad 3.9965$
Coefficients:
Estimate Std. Error t value $\operatorname{Pr}(>|\mathrm{t}|)$

| (Intercept) | 4.56481 | 0.33165 | 13.76 | $<2 \mathrm{e}-16$ *** |
| :--- | :--- | :--- | :--- | :--- |
| x | 0.50645 | 0.01154 | 43.89 | $<2 \mathrm{e}-16$ *** |

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 '*’ 0.05 '.’ 0.1 ' ' 1
Residual standard error: 1.816 on 48 degrees of freedom
Multiple R-squared: 0.9757, Adjusted R-squared: 0.9752
F-statistic: 1926 on 1 and 48 DF , p-value: < $2.2 \mathrm{e}-16$

## Illustration with R (heteroscedastic data)



## Illustration with R (heteroscedastic data)

```
reg <- lm(data=df[df$type=="heteroscedastic",],y~x)
summary(reg)
```


## Illustration with R (heteroscedastic data)

```
reg <- lm(data=df[df$type=="heteroscedastic",],y~x)
summary(reg)
```

```
Call:
lm(formula = y ~ x, data = df[df$type == "heteroscedastic", ])
```

Residuals:

| Min | $1 Q$ | Median | 3Q | Max |
| ---: | ---: | ---: | ---: | ---: |
| -25.063 | -3.472 | 0.663 | 3.707 | 19.327 |

Coefficients:
Estimate Std. Error t value $\operatorname{Pr}(>|t|)$

| (Intercept) | 4.98800 | 1.41061 | 3.536 | 0.000911 | $* * *$ |
| :--- | :--- | :--- | ---: | :--- | :--- |
| x | 0.56002 | 0.04908 | 11.411 | $2.83 \mathrm{e}-15$ | *** |

Signif. codes: 0 ‘***’ 0.001 ‘**’ 0.01 ‘*’ 0.05 '.' 0.1 ' ’ 1
Residual standard error: 7.722 on 48 degrees of freedom Multiple R-squared: 0.7306 , Adjusted R-squared: 0.725
F-statistic: 130.2 on 1 and 48 DF , p-value: 2.83e-15

## Illustration with R (quadratic data)



## Illustration with R (quadratic data)

```
reg <- lm(data=df[df$type=="quadratic",],y~x)
summary(reg)
```


## Illustration with R (quadratic data)

```
reg <- lm(data=df[df$type=="quadratic",],y~}x
summary(reg)
```

```
Call:
lm(formula = y ~ x, data = df[df$type == "quadratic", ])
```

Residuals:

| Min | $1 Q$ | Median | 3Q | Max |
| ---: | ---: | ---: | ---: | ---: |
| -11.759 | -5.847 | -2.227 | 3.746 | 17.346 |

Coefficients:
Estimate Std. Error $t$ value $\operatorname{Pr}(>|t|)$
(Intercept) $7.05330 \quad 1.41238 \quad 4.9948 .23 e-06 * * *$
$\mathrm{x} \quad 0.65517 \quad 0.0491413 .333<2 e-16$ ***
Signif. codes: $0{ }^{\prime} * * * ’ 0.001$ '**' 0.01 '*’ 0.05 '.' 0.1 ' , 1
Residual standard error: 7.732 on 48 degrees of freedom
Multiple R-squared: 0.7874 , Adjusted R-squared: 0.783
F-statistic: 177.8 on 1 and 48 DF , p-value: < 2.2e-16

## Illustration with R (quadratic data, polynomial regression)



## Illustration with R (quadratic data, polynomial regression)

```
df$x2=df$x^2
reg_quad <- lm(data=df[df$type=="quadratic",],y~x+x2)
summary(reg_quad)
```


## Illustration with R (quadratic data, polynomial regression)

```
df$x2=df$x^2
reg_quad <- lm(data=df[df$type=="quadratic",],y~x+x2)
summary(reg_quad)
```

```
Call:
lm(formula = y ~ x + x2, data = df[df$type == "quadratic", ])
```

Residuals:
Min 1Q Median 3Q Max
$-4.7834-0.8638-0.0480 \quad 1.1312 \quad 3.9913$

Coefficients:
Estimate Std. Error $t$ value $\operatorname{Pr}(>|t|)$

| (Intercept) | 5.3065389 | 0.3348067 | 15.850 | $<2 \mathrm{e}-16 * * *$ |
| :--- | :--- | :--- | ---: | ---: |
| x | 0.0036098 | 0.0252807 | 0.143 | 0.887 |
| x 2 | 0.0164635 | 0.0005694 | 28.913 | $<2 \mathrm{e}-16 * * *$ |

Signif. codes: $0{ }^{\prime} * * * ’ 0.001$ '**' 0.01 '*’ 0.05 '.' 0.1 ' , 1

Residual standard error: 1.803 on 47 degrees of freedom Multiple R-squared: 0.9887, Adjusted R-squared: 0.9882
F-statistic: 2053 on 2 and 47 DF, p-value: < $2.2 e-16$

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## Making Predictions

```
xv <- seq(-20,60,.5)
yv <- predict(reg_quad,list(x=xv,x2=xv^2))
ggplot(data=df[df$type=="quadratic",]) + theme_bw() +
    geom_hline(yintercept=0) + geom_vline(xintercept=0) +
    geom_point(aes(x=x,y=y),color="blue") +
    geom_line(data=data.frame(x=xv, y=yv), aes (x=x,y=y), color="red", siz
```



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## Confidence interval

Remember that

$$
\begin{aligned}
& \hat{\beta}=\frac{\sum_{i=1}^{n} x_{i} y_{i}-\frac{1}{n} \sum_{i=1}^{n} x_{i} \sum_{j=1}^{n} y_{j}}{\sum_{i=1}^{n}\left(x_{i}^{2}\right)-\frac{1}{n}\left(\sum_{i=1}^{n} x_{i}\right)^{2}} \\
& \hat{\alpha}=\bar{y}-\hat{\beta} \bar{x}
\end{aligned}
$$

$\hat{\beta}$ and $\hat{\alpha}$ are sums of the $\varepsilon_{i}$ 's and it is thus possible to compute confidence intervals assuming:

- the linear model holds true
- either the errors in the regression are normally distributed
- or the number of observations is sufficiently large so that the actual distribution of the estimators can be approximated using the central limit theorem


## Illustration with R

The Anscombe quartet
head $(\mathrm{a}, 10)$

|  | idx | set | x | $y$ |
| :--- | ---: | ---: | ---: | ---: |
| 1 | 1 | 1 | 10 | 8.04 |
| 2 | 1 | 2 | 10 | 9.14 |
| 3 | 1 | 3 | 10 | 7.46 |
| 4 | 1 | 4 | 8 | 6.58 |
| 5 | 2 | 1 | 8 | 6.95 |
| 6 | 2 | 2 | 8 | 8.14 |
| 7 | 2 | 3 | 8 | 6.77 |
| 8 | 2 | 4 | 8 | 5.76 |
| 9 | 3 | 1 | 13 | 7.58 |
| 10 | 3 | 2 | 13 | 8.74 |

## Illustration with R

## The Anscombe quartet

head $(a, 10)$
Confidence intervals with ggplot

```
ggplot(data=a,aes(x=x,y=y)) + theme_bw() +
    facet_wrap(~set,nrow=1) + geom_point(color="blue") +
    geom_smooth(method='lm',color="red")
```



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## Conclusion

(1) You need a model to perform your regression
(2) You need to check whether the underlying hypothesis of this model are reasonable or not

This model will allow you to:
(1) Assess and quantify the effect of parameters on the response
(2) Extrapolate within the range of parameters you tried
(3) Detect outstanding points (those with a high residual and/or with a high lever)
This model will guide on how to design your experiments:

- e.g., the linear model assumes some uniformity of interest over the parameter space range
- if your system is heteroscedastic, you will have to perform more measurements for parameters that lead to higher variance

